

Section 9: Detection v.s. Recovery^a

- Sections: Wed, 7:30-8:30pm (SC 705); OHs: Wed 8:30-9:30pm (SC 316.07).
- All the section materials (handouts & solutions) can be found either on Canvas or here.

^aThis handout is based on [9].

Definition 1. Let distributions P_n, Q_n be defined on the measurable space $(\Omega_n, \mathcal{F}_n)$. We say that the sequence Q_n is contiguous to P_n , and write $Q_n \triangleleft P_n$, if for any sequence A_n of events,

$$\lim_{n \rightarrow \infty} P_n(A_n) = 0 \implies \lim_{n \rightarrow \infty} Q_n(A_n) = 0.$$

Lemma 1. If $Q_n \triangleleft P_n$, then there is no hypothesis test of the alternative Q_n against the null P_n with $\Pr[\text{type I error}] + \Pr[\text{type II error}] = o(1)$.

Note that $Q_n \triangleleft P_n$ and $P_n \triangleleft Q_n$ are not equivalent, but either of them implies non-distinguishability.

Our goal today is to show thresholds below which spiked and unspiked random matrix models are contiguous.

Lemma 2. Let $\{P_n\}$ and $\{Q_n\}$ be two sequences of distributions on $(\Omega_n, \mathcal{F}_n)$. If the second moment

$$\mathbb{E}_{P_n} \left[\left(\frac{dQ_n}{dP_n} \right)^2 \right]$$

exists and remains bounded as $n \rightarrow \infty$, then $Q_n \triangleleft P_n$.

1. Prove this lemma.

Lemma 3. Consider a hypothesis test of a simple alternative Q against a simple null P . Let α be the probability of type I error, and β the probability of type II error. Regardless of the test, we must have

$$\frac{(1 - \beta)^2}{\alpha} + \frac{\beta^2}{(1 - \alpha)} \leq \mathbb{E}_P \left(\frac{dQ}{dP} \right)^2,$$

assuming the right-hand side is defined and finite. Furthermore, this bound is tight: for any $\alpha, \beta \in (0, 1)$ there exist P, Q and a test for which equality holds.

2. Prove the lemma above and discuss the difference between Lemma 2 and Lemma 3.

Definition 2 (Gaussian Wigner Spiked Matrix Model). *We observe $Y = \lambda xx + \frac{1}{\sqrt{n}}W$, where W is an $n \times n$ random symmetric matrix with entries drawn iid (up to symmetry) from a fixed distribution of mean 0 and variance 1.*

Question 1. *Can we “detect” whether there is a spike or not?*

3. Try to formalize the question above. Is there a difference between “detection” and “recovery”?

We will adopt a Bayesian point of view from now on. Namely, we assume a priori $x \sim \mathcal{X}$, where $\mathcal{X} = \mathcal{X}_n$ is a sequence of distributions on \mathbb{R}^n , with the default example being $\mathcal{N}(0, I_n/n)$. It is understood that $\|x\| \approx 1$.

Lemma 4. *Let $\lambda \geq 0$. Let $Q_n = \text{GWig}_n(\lambda, \mathcal{X})$ and $P_n = \text{GWig}_n(0)$. Let x and x' be independently drawn from \mathcal{X}_n . Then*

$$\mathbb{E}_{P_n} \left(\frac{dQ_n}{dP_n} \right)^2 = \mathbb{E}_{x, x'} \exp \left(\frac{n\lambda^2}{2} \langle x, x' \rangle^2 \right)$$

4. Prove Lemma 4.

It is well known that our spiked Wigner model admits the following spectral behavior.

Theorem 1. *Let Y be drawn from $\text{GWig}(\lambda, \mathcal{X})$ with any spike prior \mathcal{X} supported on unit vectors ($\|x\| = 1$):*

- *If $\lambda \leq 1$, the top eigenvalue of Y converges almost surely to 2 as $n \rightarrow \infty$, and the top (unit-norm) eigenvector v has trivial correlation with the spike: $\langle v, x \rangle^2 \rightarrow 0$ almost surely.*
- *If $\lambda > 1$, the top eigenvalue converges almost surely to $\lambda + 1/\lambda > 2$, and v estimates the spike nontrivially: $\langle v, x \rangle^2 \rightarrow 1 - 1/\lambda^2$ almost surely.*

5. Prove that for $\lambda < 1$ “detection” is impossible, assuming $x_i \stackrel{iid}{\sim} \mathcal{N}(0, 1/n)$.

6. Compare this result to Theorem 1. Do the thresholds for detection and recovery match? What about more general noise distributions and more general priors on x ?

References

- [1] BAYATI, M., AND MONTANARI, A. The dynamics of message passing on dense graphs, with applications to compressed sensing. *IEEE Transactions on Information Theory* 57, 2 (2011), 764–785.
- [2] BOLTHAUSEN, E. An iterative construction of solutions of the tap equations for the sherrington–kirkpatrick model. *Communications in Mathematical Physics* 325, 1 (2014), 333–366.
- [3] CHATTERJEE, S. A simple invariance theorem. *arXiv preprint math/0508213* (2005).
- [4] FENG, O. Y., VENKATARAMANAN, R., RUSH, C., AND SAMWORTH, R. J. A unifying tutorial on approximate message passing. *Foundations and Trends in Machine Learning* 15, 4 (2022), 335–536.
- [5] FOX, C. W., AND ROBERTS, S. J. A tutorial on variational bayesian inference. *Artificial intelligence review* 38 (2012), 85–95.
- [6] GUERRA, F., AND TONINELLI, F. L. The thermodynamic limit in mean field spin glass models. *Communications in Mathematical Physics* 230 (2002), 71–79.
- [7] MONTANARI, A., AND SEN, S. A short tutorial on mean-field spin glass techniques for non-physicists. *arXiv preprint arXiv:2204.02909* (2022).
- [8] PANCHENKO, D. *The sherrington-kirkpatrick model*. Springer Science & Business Media, 2013.
- [9] PERRY, A., WEIN, A. S., BANDEIRA, A. S., AND MOITRA, A. Optimality and sub-optimality of pca for spiked random matrices and synchronization. *arXiv preprint arXiv:1609.05573* (2016).
- [10] TALAGRAND, M. The generalized parisi formula. *Comptes Rendus Mathematique* 337, 2 (2003), 111–114.